

Extension of Hajek-Renyi inequality for negative dependence random variables

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Abstract: *In this paper we extended the Hajek-Renyi inequality, and obtain strong laws of large numbers for negative dependence (ND) random variables by use in this inequality and martingale techniques.*

Key Words: *Hajek-Renyi inequality Negative Dependent, Strong Law of Large Numbers, martingale.*

1. Introduction

Let $\{X_n, n \geq 1\}$ be a sequence of random variable's defined on a probability space (Ω, \mathcal{F}, p) . Ebrahimi and Ghosh [4] introduced the following definition.

Definition 1. The random variables X_1, \dots, X_n are said to be ND if we have

$$P\left[\bigcap_{j=1}^n (X_j \leq x_j)\right] \leq \prod_{j=1}^n P[X_j \leq x_j], \quad (1)$$

and

$$P\left[\bigcap_{j=1}^n (X_j > x_j)\right] \leq \prod_{j=1}^n P[X_j > x_j], \quad (2)$$

for all $x_1, \dots, x_n \in R$. An infinite sequence $\{X_n, n \geq 1\}$ is said to be ND if every finite subset X_1, \dots, X_n is ND. The conditions (1) and (2) are equivalent for $n = 2$, but these do not agree for $n \geq 3$ (see Bozorgnia, et. all [3])

Let $S_n = \sum_{i=1}^n X_i$. The problem convergence with probability one of $\frac{S_n - ES_n}{n}$ for sequence of negative dependence random variable's was studied by Bozorgnia et.all [3]. In addition strong laws of large numbers and complete convergence

for ND random variable's were studied by Amini and Bozorgnia [1],[2]. We will proved strong laws of large numbers for ND random variables by use Hajek-Renyi inequality. The Hajek-Renyi (1955) prove the following important inequality. If $\{X_n, n \geq 1\}$ is a sequence of independent random variables with $E(X_n) = 0$, $E(X_n^2) < \infty$, $n \geq 1$ and $\{b_n, n \geq 1\}$ is a positive nondecreasing sequence of real numbers. Then for every $\varepsilon > 0$, and positive $m < n$,

$$P\left[\max_{m \leq j \leq n} \frac{|S_j|}{b_j} \geq \varepsilon\right] \leq \frac{1}{\varepsilon^2} \left(\sum_{j=m+1}^n \frac{\sigma_j^2}{b_j^2} + \sum_{j=1}^m \frac{\sigma_j^2}{b_m^2} \right), \quad (3)$$

where $\sigma_j^2 = Var(X_j)$.

This inequality has been studied by many out-hors, the latest literature is given by J.Liu, S.Gan and P.Chen [7] for negative association random variables. In this paper, we extend this inequality for ND random variables by martingale technique and using it we proved some strong limit theorems. We need to following lemmas for next

sections.

Lemma 1(Kronecker) If $\{a_n\}$, $\{b_n\}$ are sequences of real numbers with,

$b_n \uparrow \infty$, if $\sum_{j=1}^{\infty} \frac{a_j}{b_j}$ converging, then

$$\frac{1}{b_n} \sum_{j=1}^n a_j \rightarrow 0.$$

Lemma 2([5]) The sequence $\{X_n, n \geq 1\}$ converges to X with probability one if and only if for every $\varepsilon > 0$,

$$\lim_{n \rightarrow \infty} P[\sup_{k \geq n} |X_k - X| > \varepsilon] = 0.$$

2. Extension of Hajek-Renyi inequality

In this section, we extend Hajek-Renyi inequality for ND random variables, where $E[X_n | \mathcal{F}_{n-1}] = 0$, $\mathcal{F}_n = \sigma(X_1, \dots, X_n)$ for every $n \geq 1$.

Lemma 3([6]) Let $\{X_n, \mathcal{F}_n, n \geq 1\}$ be a martingale and $\{b_n, n \geq 1\}$ be a sequence of positive nondecreasing real numbers, then for every $\varepsilon > 0$,

$$P[\max_{1 \leq k \leq n} \frac{|X_k|}{b_k} > \varepsilon] \leq \frac{1}{\varepsilon^1} (b_1^{-1} EX_1^+ + \sum_{k=2}^n b_k^{-1} (EX_k^+ - EX_{k-1}^+)), \quad (4)$$

Remark Under the assumption of lemma 3 for every $1 \leq m \leq n$, we have,

$$\begin{aligned} P[\max_{m \leq k \leq n} \frac{|X_k|}{b_k} > \varepsilon] &= P[\max_{1 \leq j \leq n-m+1} \frac{|X_{j+m-1}|}{b_{j+m-1}} > \varepsilon] \\ &\leq \frac{1}{\varepsilon^1} (b_m^{-1} EX_m^+ + \sum_{j=2}^{n-m+1} b_{j+m-1}^{-1} (EX_{j+m-1}^+ - EX_{j+m-2}^+)) \quad \text{i)} \\ &= \frac{1}{\varepsilon^1} (b_m^{-1} EX_m^+ + \sum_{k=m+1}^n b_k^{-1} (EX_k^+ - EX_{k-1}^+)) \end{aligned}$$

Theorem 1. Let $\{X_n, n \geq 1\}$ be a sequence of ND random variables with $E[X_n | \mathcal{F}_{n-1}] = 0$ and $EX_n^2 < \infty, n \geq 1$ and $\{b_n, n \geq 1\}$ be a sequence of positive nondecreasing real numbers, then for every $\varepsilon > 0$,

$$\text{i)} \quad P[\max_{1 \leq k \leq n} \frac{|S_k|}{b_k} \geq \varepsilon] \leq \frac{1}{\varepsilon^2} \sum_{j=1}^n \frac{\sigma_j^2}{b_j^2}$$

$$\text{ii)} \quad P[\max_{m \leq k \leq n} \frac{|S_k|}{b_k} \geq \varepsilon] \leq \frac{1}{\varepsilon^2} \left(\sum_{k=m+1}^n \frac{\sigma_k^2}{b_k^2} + \sum_{k=1}^m \frac{\sigma_k^2}{b_m^2} \right)$$

Proof. Since $E[X_n | \mathcal{F}_{n-1}] = 0$, then $\{S_n, \mathcal{F}_n, n \geq 1\}$ is a martingale and hence $\{|S_n|, \mathcal{F}_n, n \geq 1\}$ is a submartingale. In addition if h is an increasing function then $\{h(|S_n|), \mathcal{F}_n, n \geq 1\}$ is a submartingale, thus by lemma 3 for every $\varepsilon > 0$, we have

$$\begin{aligned} \text{i)} \quad P[\max_{1 \leq k \leq n} \frac{|S_k|}{b_k} \geq \varepsilon] &\leq P[\max_{1 \leq k \leq n} \frac{S_k^2}{b_k^2} \geq \varepsilon^2] \\ &\leq \varepsilon^{-2} [b_1^{-2} EX_1^2 + \sum_{k=2}^n b_k^{-2} (ES_k^2 - ES_{k-1}^2)] \leq \\ &\quad \frac{1}{\varepsilon^2} \sum_{j=1}^n \frac{\sigma_j^2}{b_j^2} \end{aligned}$$

ii) By above remark for $1 \leq m < n$, we obtain

$$\begin{aligned} P[\max_{m \leq k \leq n} \frac{|S_k|}{b_k} \geq \varepsilon] &\leq P[\max_{m \leq k \leq n} \frac{S_k^2}{b_k^2} \geq \varepsilon^2] \\ &\leq \frac{1}{\varepsilon^2} [b_m^{-2} ES_m^2 + \sum_{k=m+1}^n b_k^{-2} (ES_k^2 - ES_{k-1}^2)] \leq \\ &\quad \frac{1}{\varepsilon^2} \left(\sum_{k=1}^m \frac{\sigma_k^2}{b_m^2} + \sum_{k=m+1}^n \frac{\sigma_k^2}{b_k^2} \right). \end{aligned}$$

Hence complete the proof. \square .

Corollary 1. Under the assumption of theorem 1 we have

$$P[\max_{1 \leq k \leq n} \frac{|S_k|}{b_n} \geq \varepsilon] \leq \frac{1}{\varepsilon^2} \sum_{j=1}^n \frac{\sigma_j^2}{b_j^2}$$

ii)

$$P[\max_{m \leq k \leq n} \frac{|S_k|}{b_n} \geq \varepsilon] \leq \frac{1}{\varepsilon^2} \left(\sum_{k=m+1}^n \frac{\sigma_k^2}{b_k^2} + \sum_{k=1}^m \frac{\sigma_k^2}{b_m^2} \right)$$

3. Some Strong Convergence Theorems

In this section, by using Theorem 1 we can prove the following theorem for negative dependence random variable's .

Theorem 2. Let $\{X_n, n \geq 1\}$ be a sequence of ND random variables with $E[X_n | \mathcal{F}_{n-1}] = 0, n \geq 1$ and $\{b_n, n \geq 1\}$ be a sequence of positive nondecreasing real numbers. If $\sum_{n=1}^{\infty} \frac{\sigma_n^2}{b_n^2} < \infty$, then for every $0 < \beta < 2$

i) $E(\sup_n (\frac{|S_n|}{b_n})^\beta) < \infty.$

ii) If $b_n \rightarrow \infty$, as $n \rightarrow \infty$, then

$$\frac{S_n}{b_n} \rightarrow 0, \quad W.P.1.$$

Proof.

i) We have

$$E(\sup_n (\frac{|S_n|}{b_n})^\beta) < \infty \Leftrightarrow \int_1^\infty P[\sup_n \frac{|S_n|}{b_n} > t^{\frac{1}{\beta}}] dt < \infty.$$

Now by lemma 2 and Hajek-Renyi inequality we obtain

$$P[\sup_n \frac{|S_n|}{b_n} > t^{\frac{1}{\beta}}] = \lim_{m \rightarrow \infty} P[\max_{1 \leq n \leq m} \frac{|S_n|}{b_n} > t^{\frac{1}{\beta}}] \leq \frac{1}{t^{\frac{2}{\beta}}} \sum_{n=1}^\infty \frac{\sigma_n^2}{b_n^2}.$$

Hence

$$\int_1^\infty P[\sup_n \frac{|S_n|}{b_n} > t^{\frac{1}{\beta}}] dt \leq \int_1^\infty t^{-\frac{2}{\beta}} \sum_{n=1}^\infty \frac{\sigma_n^2}{b_n^2} dt = O(1) \sum_{n=1}^\infty \frac{\sigma_n^2}{b_n^2} < \infty,$$

this complete the proof.

ii) By Theorem 1, for every $\varepsilon > 0$,

$$P[\sup_{m \leq k} \frac{|S_k|}{b_k} > \varepsilon] = \lim_{n \rightarrow \infty} P[\max_{m \leq k \leq n} \frac{|S_k|}{b_k} > \varepsilon] \leq \frac{1}{\varepsilon^2} (\sum_{j=1}^m \frac{\sigma_j^2}{b_m^2} + \sum_{j=m+1}^\infty \frac{\sigma_j^2}{b_j^2}),$$

now by the assumption of $\sum_{j=1}^\infty \frac{\sigma_j^2}{b_j^2} < \infty$, and Kronecker's lemma we obtain

$$\lim_{m \rightarrow \infty} P[\sup_{k \geq m} \frac{|S_k|}{b_k} > \varepsilon] = 0,$$

hence lemma 2 complete the proof. \square

Corollary 2. Under the assumption of theorem 2, if $\sup_n \sigma_n^2 < \infty$, then for every $\alpha > \frac{1}{2}$.

i)

$$P[\sup_{j \geq m} \frac{|S_j|}{j^\alpha} > \varepsilon] \leq$$

$$\frac{1}{\varepsilon^2} (\sum_{j=1}^m \frac{1}{m^{2\alpha}} + \sum_{j=m+1}^\infty \frac{1}{j^{2\alpha}}) \sup_n \sigma_n^2.$$

ii) $\frac{S_n}{n^\alpha} \rightarrow 0 \quad W.P.1.$

iii) For every $0 < \beta < 2$,

$$E(\sup_n (\frac{|S_n|}{n^\alpha})^\beta) < \infty.$$

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